

Do Scientific Products Contribute to the Level of Output? Empirical Evidence from the European Union Countries

Çiğdem Börke TUNALI¹

Abstract

The relationship between scientific products and the level of productivity or output has become a hotly debated issue in recent years. However, the number of existing studies which examine this relationship is very low and hence more analyses are needed in order to reach conclusive results. This paper investigates the effect of scientific products on the level of gross domestic product (GDP) per capita for 15 European Union countries over the period 1981-2011 by using Autoregressive Distributed Lag (ARDL) methodology to cointegration developed by Pesaran (1997), Pesaran and Smith (1998), Pesaran and Shin (1999) and Pesaran et al. (2001). According to the results of the empirical analysis, scientific publications have a statistically significant effect on the level of GDP per capita only in France and Sweden. In France the effect of scientific publications on the level of GDP per capita is negative in the short-run and this effect becomes positive in the long-run. However, in Sweden scientific publications have a positive impact on the level of GDP per capita both in the short-run and the long-run.

JEL classification numbers: C22, O30, O31, O47, O52.

Keywords: Scientific and technical journal articles, GDP per capita, European Union countries, ARDL methodology.

¹ Assistant Professor, Department of Economics, Faculty of Economics, Istanbul University, Merkez Kampus, Beyazit, 34126, Istanbul, Turkey.

1 Introduction

The accumulation of knowledge is one of the main elements that determine the productivity and international competitiveness of a country (Inglesi-Lots and Pouris, 2013: 129). Theoretically, the significance of knowledge has been emphasized in the literature since the 1980s (Romer, 1986; Lucas, 1988). In parallel with these theoretical developments, the relationship between the knowledge stock of a country and its productivity or output level has been investigated by many researchers.

In the existing empirical literature, although there are quite a few studies which examine the relationship between the knowledge stock and productivity or output level by drawing on input indicators such as research and development expenditures the number of analyses which evaluate this relationship by using output indicators such as scientific publications is very low. Moreover, most of the existing studies which employ scientific publications to represent the knowledge stock of a country either calculate a basic correlation coefficient or apply causality analysis. Unlike most of the previous analyses, this study investigates the relationship between scientific publications and the output level by using ARDL methodology to cointegration developed by Pesaran (1997), Pesaran and Smith (1998), Pesaran and Shin (1999) and Pesaran et al. (2001) for the 15 European Union countries over the period 1981-2011. To the best of our knowledge, this is the first empirical study which analyses the linkage between scientific publications and the level of GDP per capita in the 15 European Union countries by applying the ARDL methodology to cointegration. Thus, the contribution of this study to the existing literature is two fold: First, the effect of scientific publications on the level of output is examined by using the most comprehensive data set available for the European Union countries; second, unlike most of the earlier analyses ARDL methodology to cointegration developed by Pesaran (1997), Pesaran and Smith (1998), Pesaran and Shin (1999) and Pesaran et al. (2001) is employed to shed light on the short-run and the long-run impacts of scientific publications on the level of output.

The remainder of the paper is as follows: in section 2 the existing empirical literature which focuses on the scientific publications-output level or economic growth nexus is summarized, in section 3 the model which provides the basis for the empirical analysis is discussed, in section 4 empirical methodology and the data is presented, in section 5 the results of the empirical analysis are explained and finally in section 6, the findings of the empirical analysis and policy suggestions are discussed.

2 Literature Review

In the existing literature, although there are many studies which use input indicators such as research and development expenditure to investigate the relationship between knowledge and economic growth the number of studies which draw on output indicators of knowledge such as scientific publications is very low. Here, the results of the papers which employ output indicators of knowledge in their empirical analyses are summarized.

Moya-Anegon and Herrero-Solana (1999) analyze the relationship between scientific publications and GDP by calculating the correlation between GDP in 1995 and the number of articles in Science Citation Indexed (SCI) journals in 1996 for the Latin American countries. According to the calculations, Moya-Anegon and Herrero-Solana state that there is a positive relationship between GDP level of the countries under investigation and their scientific publications.

King (2004) investigates the relationship between wealth intensity represented by GDP per capita and citations intensity represented by the ratio of the citations to all papers to the GDP in 31 countries containing G8 group and 15 European Union countries (countries which were members of the European Union before 2004 accession) in 1995. King (2004) finds that Scandinavian countries, Israel, the Netherlands and Switzerland perform very well with regard to this measure.

Vinkler (2008) estimates the correlation between GDP and the number of publications in 14 European Community member states, the US, Japan and 10 Central and Eastern European countries from 1975 to 2004 by using synchronized and asynchronized (examining consecutive time periods) longitudinal analyses. While the results of synchronized longitudinal analyses indicate that GDP and the number of publications move together the results of consecutive studies could not support that scientific publications may lead to economic growth. Hence, Vinkler (2008) suggests that it might be hazardous to assume a simple linear or exponential relationship between GDP and scientific products and more analyses are needed to shed light on the relationship between these variables.

Lee et al. (2011) examine the relationship between research output and economic productivity in a number of Asian and Western countries over the period 1981-2007. The authors draw on Granger Causality Analysis in their empirical evaluation and find a reciprocal relationship between research output and economic productivity in Asian countries while the relationship between these variables in Western countries is ambiguous.

Meo et al. (2013) calculate the correlation between GDP per capita and research outcomes such as the number of published documents, citations per documents

and H-index in 40 Asian countries between 1996 and 2011. According to the results of this analysis, the authors argue that there is no relationship between GDP per capita and research outcomes.

Inglesi-Lots and Pouris (2013) analyze the effect of research output on economic growth in South Africa over the period 1980-2008. By using ARDL methodology the authors conclude that while the research output affects economic growth in South Africa economic growth does not have an effect on research output. Hence, the results of this empirical analysis indicate a uniletaral relationship between knowledge and economic growth.

Kim and Lee (2015) investigate the impact of scientific and technological knowledge on economic growth in East Asian and Latin American countries between 1960 and 2005. In the empirical analysis, scientific and technological knowledge is represented by the number of Science Citation Indexed (SCI) articles per million residents and the number of corporate patents granted by the United States Patent and Trademark Office (USPTO) per million residents. Kim and Lee (2015) estimate the regression models by using OLS, fixed effects and Generalized Methods of Moments (GMM) estimators and come to the conclusion that whilst the number of SCI articles per million residents does not have a statistically significant effect on economic growth the number of corporate patents has a significant and positive effect on economic growth in the countries under investigation.

Inglesi-Lotz et al. (2015) examine the relationship between economic growth and research output in the BRICS countries (Brazil, Russia, India, China and South Africa) during the period 1981-2011. In the empirical analysis, Inglesi-Lotz et al. (2015) draw on panel causality analysis and take into account both cross-sectional dependency and heterogeneity among the countries. The results of the empirical analysis indicate that there is a bidirectional relationship between economic growth and research output in India, whereas in Brazil, China and South Africa there is no relationship between these variables. Therefore, Inglesi-Lotz et al. (2015) conclude that the policies of the BRICS countries in relation to research and development are not similar although their economies show similar characteristics.

In a recent study, Kumar et al. (coming soon) assess the effect of scientific and technical research on economic growth in China and the USA over the period 1981-2012. The authors use ARDL methodology developed by Pesaran (1997), Pesaran and Smith (1998), Pesaran and Shin (1999) and Pesaran et al. (2001) and Toda and Yamamoto (1995) non-Granger causality tests. According to the results of the empirical analysis, Kumar et al. (coming soon) argue that research publications have a positive and statistically significant effect on economic growth both in the short-run and the long-run in China and the USA. Moreover, by taking

into account the results of causality analysis the authors suggest that there is a bidirectional causality between output per worker and research publications per worker in China while in the USA this relationship is unidirectional which is from the research publications per worker to the output per worker.

In summary, although the empirical analyses which focus on the relationship between scientific publications and economic growth are on the rise in recent years the number of existing studies is still very low. Moreover, most of these studies calculate either a simple correlation coefficient between scientific publications and economic growth or apply basic causality analysis. Therefore, more analyses are needed to clarify the exact relationship between scientific publications and economic growth in different countries.

3 Model

In the empirical analysis, the modelling framework which is used by Kumar et al. (coming soon) to investigate the relationship between research output and economic growth in China and the USA is applied. This modelling framework which is introduced by Sturm et al. (1998) and Rao (2010) pertains to the Augmented Solow Growth Model (Solow, 1956). In this modelling framework, production function is described as follows (Kumar et al., coming soon):

$$Y_t = A_t SP_t^\sigma K_t^\alpha L_t^\beta \quad (1)$$

In this equation Y represents GDP level, A is the stock of technology, SP is the number of scientific publications, K is the capital stock, L is the labor stock, σ , α and β is the share of scientific publications, the share of capital stock and the share of labor stock in the production function respectively and t is the time subscript. If this equation is divided by the labor stock the following equation is obtained:

$$y_t = A_t sp_t^\sigma k_t^\alpha \quad (2)$$

where y, sp and k is the GDP, the number of scientific publications and capital per worker respectively. In order to have a testable equation Kumar et al. (coming soon) assume $A_t = A_0 e^{g^T}$ and get:

$$y_t = A_0 e^{g^T} sp_t^\sigma k_t^\alpha \quad (3)$$

Kumar et al. (coming soon) attain the equation below which is estimated in the empirical analysis by taking the logarithm of equation 3:

$$\ln y_t = \theta + \alpha \ln k_t + \sigma \ln sp_t + \varepsilon_t \quad (4)$$

where θ is the constant term, α is the share of capital, σ is the share of scientific publications which measures the production elasticity of this variable and ε is the error term.

4 Empirical Methodology and Data

Similar to the Kumar et al. (coming soon)'s study, the Autoregressive Distributed Lag (ARDL) approach to cointegration introduced by Pesaran (1997), Pesaran and Smith (1998), Pesaran and Shin (1999) and Pesaran et al. (2001) is employed in order to estimate equation 4. As it is very well known, there are a number of different methods such as residual based test of Engle and Granger (1987) and the maximum likelihood based tests of Johansen (1991, 1995) and Johansen and Juselius (1990) for cointegration analysis in the existing literature. However, the ARDL methodology to cointegration (Pesaran et al., 2001) has several advantages in comparison to the other methods. First, ARDL methodology can be applied to the variables which are stationary (i.e. I(0)), integrated of order one (i.e. I(1)) or integrated fractionally (Ghosh, 2009: 700). Second, ARDL methodology circumvents the problems of serial correlation and endogeneity by proper augmentation (Ghatak and Siddiki, 2001: 574). Third, by applying a simple linear transformation the Error Correction Model (ECM) which combines short run adjustments and long-run equilibrium can be obtained from the ARDL model (Jalil and Mahmud, 2009: 5168, 5169). Last but not least, ARDL methodology provides more reliable results than the other methods when the sample size is small (Pesaran and Shin, 1999). The last characteristics of the ARDL methodology is especially important in this empirical analysis since the data for the scientific publications is not available for a long period of time.

By taking into account equation 4, the ARDL model which is estimated can be stated as follows:

$$\begin{aligned} \Delta \ln y_t = & \alpha_{10} + \alpha_{11} \ln y_{t-1} + \alpha_{12} \ln k_{t-1} + \alpha_{13} \ln sp_{t-1} + \alpha_{14} crisis_t + \alpha_{15} trend_t + \\ & \sum_{i=1}^m \beta_{11i} \Delta \ln y_{t-i} + \sum_{i=0}^m \beta_{12i} \Delta \ln K_{t-i} + \sum_{i=0}^m \Delta \ln sp_{t-i} + \varepsilon_{1t} \end{aligned} \quad (5)$$

Equation 5 has the same variables with equation 4 and two additional variables. The first additional variable is the crisis which represents 2008 Global Economic Crisis and the second additional variable is the trend term which represents the time trend in the regressions.

The steps of the ARDL model are as follows (Jalil and Mahmud, 2009: 5169): At first, equation 5 is estimated by using Ordinary Least Squares (OLS) estimator and then, F-test is calculated to test whether there is a long-run relationship between the variables. The null hypothesis of this test is $H_0: \alpha_{11}=\alpha_{12}=\alpha_{13}=\alpha_{14}=\alpha_{15}=0$ which means that there is no long-run relationship between the variables whilst the alternative is $H_1: \alpha_{11}\neq 0, \alpha_{12}\neq 0, \alpha_{13}\neq 0, \alpha_{14}\neq 0, \alpha_{15}\neq 0$. The calculated F-statistics is compared with two critical values. If the F-statistics is higher than the upper value then the null hypothesis of no cointegration is rejected. Instead of this, if the F-statistics drops below the lower value then the null hypothesis cannot be rejected. Finally, if the F-statistics is between the lower and higher bounds then the result is inconclusive.

The F-statistics and the lower and higher critical values are based on the size of the sample, the model specification and the lag-length (Kumar et al., coming soon). Although there are some studies which provide critical values such as Pesaran et al. (2001) and Narayan (2005) these critical values are calculated for specific sample sizes and do not take into account structural breaks (Kumar et al., coming soon). There is a built in function in Microfit 5.0 developed by Pesaran and Pesaran (2009) which calculates F-statistics and the lower and upper critical values according to the features of the sample under investigation. This built in function in Microfit 5.0 is used in this empirical analysis.

As stated above, the ARDL model can be applied as long as the variables are either I(0) or I(1). However, Ouattara (2004) claims that if the variables are integrated of order two or higher then the F-statistics provided by Pesaran et al. (2001) is not valid. Thus, in order to ensure that the variables are either I(0) or I(1) the generalised least squares (GLS) detrended version of the conventional Dickey-Fuller test (DFGLS or ERS test) proposed by Elliott et al. (1996) and a modified version of the tests of Phillips and Perron (1988) proposed by Ng and Perron (2001)² are applied before the ARDL estimation. Both of these tests perform better than the traditional unit root tests especially when the length of the series is short (Elliott et al., 1996; Ng and Perron, 2001), as is the case with this empirical analysis. After the unit root analysis, the ARDL methodology is applied to the countries whose variables are integrated of order zero or one.

² The modified Ng and Perron test is the M^{GLS} extension of the M tests developed by Perron and Ng (1996) that allows for GLS detrending.

The data covers 15 European Union countries (Austria, Belgium, Denmark, Finland, France, Germany, Greece, Ireland, Italy, Luxembourg, the Netherlands, Portugal, Spain, Sweden and the United Kingdom) over the period 1981-2011. The time period and the countries are chosen according to the data availability. Since labor force statistics are not available for most of the countries over the whole time period under investigation population statistics are used while calculating per capita series. GDP, net capital stock and population data are taken from the Annual Macro-economic Database (AMECO) of the European Commission's Directorate General for Economic and Financial Affairs (2015). Both GDP and net capital stock series are at 2010 prices. The data for the number of scientific and technical journal articles are taken from the World Bank-World Development Indicators Database (2015). The scientific and technical journal articles cover the fields of physics, biology, chemistry, mathematics, clinical medicine, biomedical research, engineering and technology, and earth and space sciences and count the articles published in the journals classified by the Institute for Scientific Information's Science Citation Index (SCI) and Social Science Citation Index (SSCI) (World Bank, 2015). Since the number of scientific and technical articles data are missing for the years of 1982, 1983 and 1984 linear interpolation method is used in order to fill in these missing values. All variables (GDP, net capital stock and the number of scientific and technical articles) are stated as a percentage of total population and their natural logarithms are taken before the estimations.

5 Empirical Results

Table 1 presents the results of the unit root tests. As it is clearly seen, except Ireland and Spain all series are either integrated of the order zero or one for all the countries under investigation. So, ARDL methodology can be applied to the 13 European Union countries excluding Ireland and Spain³.

³ As it is very well known, the presence of structural breaks in the series may cause misleading results in the unit root analysis. As demonstrated by Perron (1989) when the structural breaks were taken into account many time series regarded as non-stationary according to the conventional unit root tests were in fact stationary. Because of this reason, Clemente, Montanes and Reyes (1998) unit root test which takes into account structural breaks in the series is also applied. The results of this test are similar to the DFGLS and Ng-Perron unit root test results. These results can be provided upon request.

Table 1: Unit Root Test Results

Countries	Variables	DFGLS Test	Ng-Perron Test			
		t-Statistics	MZa	MZt	MSB	MPT
Austria	GDP	-2.40837	-15.7779*	-2.65932*	0.16855**	6.63055*
	Cap. Stock	-1.46620	-352.148***	-13.2164***	0.03753***	0.36833***
	Sci. Papers	-1.848055	-20.7196**	-3.16308**	0.15266**	4.72792**
	Δ GDP	-4.379364***	-14.0038***	-2.64069***	0.18857**	1.77004***
	Δ Cap. Stock	-0.798482	---	---	---	---
	Δ Sci. Papers	-1.912686*	---	---	---	---
Belgium	GDP	-1.729137	-10.2026	-2.01664	0.19766	9.96591
	Cap. Stock	-1.224949	-14.8788*	-2.51191*	0.16883*	7.32530*
	Sci. Papers	-1.384990	-4.33728	-1.27880	0.29484	19.2938
	Δ GDP	-4.154756***	-13.6012**	-2.59990***	0.19115**	1.83133**
	Δ Cap. Stock	-1.900395*	-9.62383**	-2.02530**	0.21045**	3.16861**
	Δ Sci. Papers	-5.238841***	-14.4577***	-2.67467***	0.18500**	1.74723***
Denmark	GDP	-1.787906	-11.3307	-2.13554	0.18847	9.21042
	Cap. Stock	-2.515187	-18.4526**	-2.96661**	0.16077**	5.35776**
	Sci. Papers	-3.103148*	-11.2960	-2.36090	0.20900	8.14533
	Δ GDP	-3.415115***	-11.7689**	-2.40414**	0.20428**	2.16500**
	Δ Cap. Stock	-3.079734***	---	---	---	---
	Δ Sci. Papers	-8.020864***	-11.8451**	-2.34630**	0.19808**	2.39951**
Finland	GDP	-2.840199	-18.0775**	-2.98269**	0.16499**	5.18234**
	Cap. Stock	-1.911926	-6.55501	-1.77537	0.27084	13.9066
	Sci. Papers	-1.147499	-44.7718***	-4.61989***	0.10319***	2.59401***
	Δ GDP	-3.353133***	---	---	---	---
	Δ Cap. Stock	-2.876647***	-17.7909***	-2.94586***	0.16558***	1.50879***
	Δ Sci. Papers	-2.506854**	---	---	---	---
France	GDP	-2.085413	-13.6185	-2.44711	0.17969	7.57630
	Cap. Stock	-1.333494	-5.64077	-1.50559	0.26691	15.7381
	Sci. Papers	-2.067474	-19.1592**	-3.02504**	0.15789**	5.17136**
	Δ GDP	-3.551732***	-12.3836**	-2.48834**	0.20094**	1.97842**
	Δ Cap. Stock	-2.589636**	-16.2615***	-2.75356***	0.16933***	1.86157**
	Δ Sci. Papers	-1.164957*	---	---	---	---
Germany	GDP	-1.486511	-4.09418	-1.35969	0.33210	21.4603
	Cap. Stock	-1.028283	-2.61075	-0.94499	0.36196	28.3104
	Sci. Papers	-1.663259	-5.09794	-1.52956	0.30004	17.5592
	Δ GDP	-4.634642***	-14.1069***	-2.63548***	0.18682***	1.81351**
	Δ Cap. Stock	-4.927382***	-14.4317***	-2.67815***	0.18557***	1.72810***
	Δ Sci. Papers	-4.745628***	-14.3045***	-2.67306***	0.18687**	1.71769***

Table 1: Unit Root Test Results (Continued)

Countries	Variables	DFGLS Test	Ng-Perron Test			
		t-Statistics	MZa	MZt	MSB	MPT
Greece	GDP	-3.145254*	-5.33891	-1.42658	0.26720	16.3467
	Cap. Stock	-3.282129**	-124.619***	-7.89309***	0.06334***	0.73311***
	Sci. Papers	-2.695774	-108.299***	-7.20758***	0.06655***	1.37739***
	Δ GDP	-1.025322	-14.3263***	-2.41153**	0.16833***	2.66146**
	Δ Cap.	---	---	---	---	---
	Δ Sci.	-1.740885*	---	---	---	---
Ireland	GDP	-2.696462	0.65207	0.45991	0.70531	113.798
	Cap. Stock	-2.583557	-46.2432***	-4.79676***	0.10373***	2.02933***
	Sci. Papers	-3.074583*	-8.22973	-1.98487	0.24118	11.1986
	Δ GDP	-1.422151	-27.7876***	-3.72665***	0.13411***	0.88419***
	Δ Cap.	-2.023370**	---	---	---	---
	Δ Sci.	-1.633474*	-2.52584	-1.11464	0.44129	9.65213
Italy	GDP	-1.342392	-6.11861	-1.48092	0.24203	14.6494
	Cap. Stock	-1.225113	-88.0718***	-6.44570***	0.07319***	1.76928***
	Sci. Papers	-1.574347	-15.8275*	-2.61642*	0.16531**	6.87420
	Δ GDP	-3.198008***	-11.3365**	-2.37437**	0.20945**	2.18602**
	Δ Cap.	-1.558501	---	---	---	---
	Δ Sci.	-0.127229	-15.8275*	-2.61642*	0.16531**	6.87420
Lux.bourg	GDP	-1.323114	-4.97763	-1.31500	0.26418	17.0023
	Cap. Stock	-2.395519	-9.97340	-2.20373	0.22096	9.26468
	Sci.Papers	-2.549846	-8.18066	-1.83945	0.22485	11.6390
	Δ GDP	-3.963092***	-13.3174**	-2.56547**	0.19264**	1.89673**
	Δ Cap.	-2.939558***	-10.5955**	-2.27039**	0.21428**	2.43262**
	Δ Sci.	-2.743093***	-3.61634	-1.27637	0.35295	6.77400
Net.lands	GDP	-2.469679	-14.5195	-2.53255	0.17442	7.17983
	Cap. Stock	-1.874049	-6.65938	-1.78153	0.26752	13.7000
	Sci. Papers	-2.642088	-3825.63***	-43.7334***	0.01143***	0.02535***
	Δ GDP	-2.794861***	-8.85923**	-2.10182**	0.23725**	2.77629**
	Δ Cap.	-3.693719***	-31.9427***	-3.96709***	0.12419***	0.85473***
	Δ Sci.	-1.484897	---	---	---	---
Portugal	GDP	-1.837332	-32.1459***	-3.90008***	0.12132***	3.43551***
	Cap. Stock	-0.421308	-26.0271***	-3.39580**	0.13047***	4.70056**
	Sci. Papers	-1.630129	-7.40560	-1.74435	0.23554	12.6128
	Δ GDP	-2.407225**	---	---	---	---
	Δ Cap.	-1.246638	---	---	---	---
	Δ Sci.	-3.872206***	-11.8934**	-2.43752**	0.20495**	2.06409**

Table 1: Unit Root Test Results (Continued)

Countries	Variables	DFGLS Test	Ng-Perron Test			
		t-Statistics	MZa	MZt	MSB	MPT
Spain	GDP	-2.117402	-54.9136***	-5.06974***	0.09232***	2.45007***
	Cap. Stock	-1.430744	-14.9094	-2.56125	0.17179*	7.06247
	Sci. Papers	-2.467587	-1.63865	-0.67239	0.41033	36.9749
	Δ GDP	-1.955749**	---	---	---	---
	Δ Cap. Stock	-2.675308***	-25.7279***	-3.48792***	0.13557***	1.26791***
	Δ Sci. Papers	0.045437	-0.22208	-0.19734	0.88859	43.0836
Sweden	GDP	-2.247323	-10.6927	-2.30857	0.21590	8.53955
	Cap. Stock	-1.624405	-7.17305	-1.77273	0.24714	12.8683
	Sci. Papers	-1.179542	-18.1728**	-2.87158*	0.15802**	5.84847*
	Δ GDP	-4.002636***	-13.4094**	-2.58875***	0.19306**	1.82933**
	Δ Cap. Stock	-2.258177**	-11.1624**	-2.30275**	0.20629**	2.42303**
	Δ Sci. Papers	-1.518626	---	---	---	---
UK	GDP	-2.506448	-24.7257***	-3.33251**	0.13478***	4.73766**
	Cap. Stock	-1.747057	-150.151***	-8.59425***	0.05724***	0.82403***
	Sci. Papers	-1.128176	-2.26721	-0.86244	0.38040	31.1971
	Δ GDP	-2.955662***	---	---	---	---
	Δ Cap. Stock	-0.956099	---	---	---	---
	Δ Sci. Papers	-5.976788***	-14.2737***	-2.66333***	0.18659**	1.74725***

Note: ***, **, * denotes the significance at 1%, 5% and 10% levels respectively. A constant and a trend term added to the models according to the examinations of the series' graphs.

Source: Author's estimations.

After examining the order of integration of the series the optimal lag length is selected according to the Log Likelihood (LogL), Sequential Modified LR Test Statistic (LR), Final Prediction Error (FPE), Akaike Information Criterion (AIC), Schwarz Information Criterion (SC) and Hannan-Quinn Information Criterion (HQ). Table 2 shows these statistics. According to the results, the optimal lag length is: 1 for Luxembourg, 2 for Belgium, Denmark and the United Kingdom, 3 for France, Germany and Portugal, 4 for Finland, Italy and Sweden, 2 or 3 for Austria, 1 or 4 for Greece and 2 or 4 for the Netherlands. Hence, these lags are used while estimating the ARDL model⁴.

⁴ For Austria, Greece and the Netherlands two ARDL models are estimated by using two different lag lengths selected according to the test statistics. The results of these models are very similar. Since the data set does not cover a long period of time here the results of the model which is estimated by using smaller lag length are presented.

Following the selection of optimal lag length F-statistics is calculated in order to decide whether there is a long-run relationship between the variables. Table 3 presents these results. As it is clearly seen, the long-run relationship between GDP, capital stock and scientific and technical articles per capita exists only for France, the Netherlands and Sweden. Thus, error-correction model (ECM) is estimated for these countries.

Table 4 and table 5 show the estimated long-run coefficients and the ECM respectively. According to the long-run coefficient estimates (table 4), scientific and technical journal articles are statistically significant and have a positive sign for France and Sweden. However, this variable is not statistically significant for the Netherlands. The results indicate that a 1% increase in scientific and technical journal articles per capita leads to 2.59 % and 2.34% increase in GDP per capita in France and Sweden respectively. Therefore, it can be suggested that while scientific and technical journal articles have a strong positive effect on GDP per capita in France and Sweden in the long-run this variable does not have an impact on GDP per capita in the Netherlands. When the results of the ECM are examined (table 5) it is seen that error-correction coefficient is statistically significant and has a negative sign for all of the countries under investigation. This result confirms the existence of a long-run relationship between the variables. With regard to the coefficient estimates, it can be stated that capital per capita and the crisis variables are statistically significant in all of the countries. While capital per capita has a positive effect crisis has a negative impact on GDP per capita in the short-run. When it comes to the scientific and technical journal articles per capita, it can be suggested that this variable is statistically significant in France and Sweden. In contrast to the long-run results, scientific and technical journal articles per capita has a negative effect on GDP per capita in France in the short-run. However, in Sweden the impact of scientific and technical journal articles on GDP per capita is positive. Therefore, it can be stated that although scientific products in France have a negative impact on GDP in the short-run this negative impact becomes positive in the long-run. Conversely, in Sweden scientific products have a positive effect on GDP per capita both in the short-run and the long-run. This may stem from the fact that the transformation of scientific knowledge to marketable products takes a long time in France while in Sweden scientific knowledge can be turned into commercial products even in the short-run.

Table 2: Lag Length Selection

Lag	LogL	LR	FPE	AIC	SC	HQ
Austria						
0	213.5031	---	5.31e-11	-15.14838	-14.71643	-15.01994
1	258.7484	70.38161	3.69e-12	-17.83322	-16.96933	-17.57634
2	275.6268	22.50448*	2.18e-12	-18.41680	-17.12096*	-18.03148
3	286.5597	12.14768	2.13e-12*	-18.55998	-16.83220	-18.04622*
4	296.7787	9.083554	2.44e-12	-18.65027*	-16.49055	-18.00807
Belgium						
0	205.8181	---	9.38e-11	-14.57912	-14.14717	-14.45068
1	263.6084	89.89609	2.58e-12	-18.19322	-17.32933	-17.93634
2	282.8034	25.59324*	1.28e-12*	-18.94840*	-17.65256*	-18.56308*
3	287.4038	5.111555	2.00e-12	-18.62250	-16.89472	-18.10874
4	298.2414	9.633499	2.19e-12	-18.75863	-16.59890	-18.11643
Denmark						
0	195.6290	---	1.99e-10	-13.82437	-13.39243	-13.69593
1	258.0556	97.10798	3.89e-12	-17.78190	-16.91800	-17.52502
2	277.8760	26.42722*	1.85e-12*	-18.58341*	-17.28757*	-18.19809*
3	283.8323	6.618053	2.61e-12	-18.35794	-16.63016	-17.84418
4	294.6155	9.585126	2.86e-12	-18.49004	-16.33031	-17.84784
Finland						
0	144.2952	---	8.94e-09	-10.02186	-9.589919	-9.893424
1	247.3855	160.3627	8.57e-12	-16.99152	-16.12763	-16.73464
2	266.7474	25.81595*	4.21e-12	-17.75907	-16.46323*	-17.37375
3	276.4990	10.83511	4.49e-12	-17.81474	-16.08696	-17.30098
4	292.0588	13.83090	3.46e-12*	-18.30065*	-16.14092	-17.65845*
France						
0	205.3292	---	9.72e-11	-14.54291	-14.11096	-14.41447
1	284.9822	123.9045	5.29e-13	-19.77646	-18.91257	-19.51958
2	310.5915	34.14584	1.64e-13	-21.00678	-19.71094	-20.62146
3	327.1631	18.41285*	1.05e-13*	-21.56764	-19.83985*	-21.05388*
4	337.8741	9.520850	1.16e-13	-21.69437*	-19.53465	-21.05217
Germany						
0	164.3353	---	2.03e-09	-11.50632	-11.07438	-11.37788
1	203.9147	61.56794	2.15e-10	-13.77146	-12.90757*	-13.51458
2	213.9882	13.43134	2.10e-10	-13.85098	-12.55514	-13.46566
3	231.5216	19.48155*	1.26e-10*	-14.48308	-12.75530	-13.96932*
4	241.3150	8.705232	1.48e-10	-14.54185*	-12.38212	-13.89965
Greece						
0	170.0570	---	1.33e-09	-11.93015	-11.49820	-11.80171
1	236.7876	103.8033*	1.88e-11	-16.20649	-15.34260*	-15.94961
2	249.1859	16.53097	1.55e-11*	-16.45821	-15.16237	-16.07289
3	258.9881	10.89138	1.64e-11	-16.51764	-14.78985	-16.00388
4	271.4388	11.06725	1.59e-11	-16.77324*	-14.61351	-16.13104*

Table 2: Lag Length Selection (Continued)

Lag	LogL	LR	FPE	AIC	SC	HQ
Italy						
0	-221.0748	---	5073.124	17.04258	17.47452	17.17102
1	-141.7920	123.3289	28.37287	11.83644	12.70033*	12.09332
2	-133.0017	11.72035	30.50190	11.85198	13.14782	12.23730
3	-125.0616	8.822327	37.21554	11.93049	13.65827	12.44425
4	-104.9524	17.87482*	20.46902*	11.10759*	13.26732	11.74979*
Lux.bourg						
0	136.7847	---	1.56e-08	-9.465530	-9.033585	-9.337090
1	180.9473	68.69743*	1.18e-09*	-12.07017	-11.20628*	-11.81329*
2	184.8799	5.243508	1.81e-09	-11.69481	-10.39897	-11.30949
3	194.0696	10.21076	2.02e-09	-11.70886	-9.981077	-11.19510
4	209.8640	14.03949	1.53e-09	-12.21215*	-10.05242	-11.56995
Net.lands						
0	208.9446	---	7.44e-11	-14.81071	-14.37877	-14.68227
1	267.0846	90.43994	1.99e-12	-18.45071	-17.58682	-18.19383
2	289.8676	30.37738*	7.60e-13	-19.47167	-18.17584*	-19.08635
3	302.5337	14.07348	6.53e-13*	-19.74324	-18.01546	-19.22948
4	313.9293	10.12941	6.85e-13	-19.92069*	-17.76096	-19.27849*
Portugal						
0	160.1085	---	2.77e-09	-11.19323	-10.76128	-11.06478
1	233.8108	114.6480	2.34e-11	-15.98599	-15.12210	-15.72911
2	245.1735	15.15018	2.08e-11	-16.16100	-14.86516	-15.77568
3	265.0588	22.09484*	1.05e-11*	-16.96732	-15.23954*	-16.45356*
4	274.4580	8.354834	1.27e-11	-16.99689*	-14.83716	-16.35469
Sweden						
0	179.4483	---	6.61e-10	-12.62580	-12.19386	-12.49736
1	267.4947	136.9609	1.93e-12	-18.48109	-17.61719	-18.22421
2	280.4793	17.31289	1.52e-12	-18.77625	-17.48041	-18.39093
3	299.0320	20.61413*	8.47e-13	-19.48386	-17.75607	-18.97009
4	316.1577	15.22283	5.81e-13*	-20.08576*	-17.92603*	-19.44356*
UK						
0	186.4565	---	3.94e-10	-13.14493	-12.71298	-13.01649
1	263.0786	119.1899	2.68e-12	-18.15397	-17.29008	-17.89709
2	285.6366	30.07729*	1.04e-12*	-19.15827	-17.86243*	-18.77295*
3	289.4357	4.221259	1.72e-12	-18.77302	-17.04523	-18.25926
4	306.7724	15.41035	1.16e-12	-19.39054*	-17.23082	-18.74834

Note: LogL: Log Likelihood, LR: Sequential Modified LR Test Statistic, FPE: Final Prediction Error, AIC: Akaike Information Criterion, SC: Schwarz Information Criterion, HQ: Hannan-Quinn Information Criterion

Source: Author's estimations.

Table 3: Results of Bounds Test at 5% Level

Countries	F Statistic	W Statistic	Outcome
Austria	5.5847 (6.3568, 7.4286)	16.7540 (19.070, 22.285)	No Cointegration
Belgium	2.8317(6.3568, 7.4286)	8.4951 (19.0704, 22.2857)	No Cointegration
Denmark	2.9348 (6.3568,7.4286)	8.8045 (19.0704,22.2857)	No Cointegration
Finland	5.4084 (6.4597, 7.5938)	16.2251 (19.3792, 22.7814)	No Cointegration
France	40.5977(6.4247, 7.4867)	121.7930(19.2741, 22.4602)	Cointegration
Germany	5.4339 (6.4247, 7.4867)	16.3016 (19.2741, 22.4602)	No Cointegration
Greece	2.5160(6.4597, 7.5938)	7.5479 (19.3792, 22.7814)	No Cointegration
Italy	0.1829 (6.4597, 7.5938)	0.5487 (19.3792, 22.7814)	No Cointegration
Lux.bourg	7.2336 (6.2915, 7.3503)	21.7008 (18.8744, 22.0508)	No Cointegration
Net.lands	9.7540(6.3568, 7.4286)	29.2620 (19.0704, 22.2857)	Cointegration
Portugal	5.2418(6.4247, 7.4867)	15.7255 (19.2741, 22.4602)	No Cointegration
Sweden	15.8066(6.4597, 7.5938)	47.4199(19.3792, 22.7814)	Cointegration
UK	2.5513 (6.4247, 7.4867)	7.6538(19.2741, 22.4602)	No Cointegration

Note: The lower and upper critical values are in parenthesis.

Source: Author's estimations.

Finally, the stability of the model should be checked by employing the CUSUM and CUSUMSQ technique which is developed by Brown et al. (1975). Appendix A, appendix B and appendix C present the plots of CSUM and CSUMSQ for France, the Netherlands and Sweden respectively. As it is clearly seen the plots of CSUM and CSUMSQ are within the 95% critical bounds implying that all coefficients estimated in the ECM are stable over the sample period under investigation.

Table 4: Estimated Long-run Coefficients

Variables	Coefficient	Stan. Error	T-Ratio
France			
Cap. Stock	-0.4287	0.7313	-0.5862
Sci. Papers	0.3293	0.1270	2.5933**
Crisis	-0.0414	0.0088	-4.7237***
Constant	22.2045	11.7868	1.8838*
Trend	0.0181	0.0092	1.9613*
Diagnostic Tests	LM Test Stat.	P-value	
Serial Correlation	1.0359	0.309	
Normality	1.6474	0.439	
Heteroscedasticity	4.1144	0.043	
Netherlands			
Cap. Stock	-0.0869	1.1876	-0.0732
Sci. Papers	0.0512	0.0691	0.7408
Crisis	-0.0644	0.0182	-3.5437***
Constant	16.0128	18.5762	0.8620
Trend	0.0222	0.0186	1.1939
Diagnostic Tests	LM Test Stat.	P-value	
Serial Correlation	0.032446	0.857	
Normality	6.4747	0.039	
Heteroscedasticity	2.1463	0.143	
Sweden			
Cap. Stock	-2.5472	0.7713	-3.3026***
Sci. Papers	0.5478	0.2340	2.3411**
Crisis	-0.0631	0.0189	-3.3465***
Constant	64.4651	14.6116	4.4119***
Trend	0.0614	0.0102	6.0128***
Diagnostic Tests	LM Test Stat.	P-value	
Serial Correlation	11.9393	0.001	
Normality	2.6365	0.268	
Heteroscedasticity	0.48016	0.488	

Note: ***, **, * denotes the significance at 1%, 5% and 10% levels respectively. Besides long-run coefficient estimates the table also shows the results of diagnostic tests. According to the normality test results, the disturbance terms are not normally distributed for the Netherlands. However, since the number of observations is higher than 30 this does not create a problem. Similarly, the results indicate that there is a problem of heteroscedasticity and serial correlation for France and Sweden respectively. Since ARDL model is robust against serial correlation (Laurenceson and Chai 2003: 30) and detecting heteroscedasticity is natural due to the mixed orders of integration of the series (Shrestha and Chowdhury, 2005) these results do not indicate a problem.

Source: Author's estimations.

Table 5: Error-Correction Model Results

Variables	Coefficient	Stan. Error	T-Ratio
France			
ARDL (3,1,3) selected based on Akaike Information Criterion			
Δ GDP1	-0.2760	0.1333	-2.0711*
Δ GDP2	-0.2519	0.1372	-1.8365*
Δ Cap. Stock	5.2821	0.5660	9.3318***
Δ Sci. Papers	0.0231	0.0351	0.6588
Δ Sci. Papers1	-0.1668	0.0388	-4.5778***
Δ Sci. Papers2	-0.0809	0.0388	-2.0863*
Δ Crisis	-0.0218	0.0056	-3.9291***
Δ Trend	0.0096	0.0029	3.3338***
ecm (-1)	-0.5275	0.1417	-3.7226***
R-Squared: 0.95200 R-Bar Squared: 0.91900		Akaike Info. Criterion: 108.9019	
SER: 0.0042661		Schwarz Bayesian Criterion: 100.9087	
DW-Statistic: 2.2582		F-Statistic: 35.2604 (0.000)	
The Netherlands			
ARDL (1,1,0) selected based on Akaike Information Criterion			
Δ Cap. Stock	2.9309	0.5618	5.2174***
Δ Sci. Papers	0.0202	0.0287	0.7042
Δ Crisis	-0.0254	0.0089	-2.8609***
Δ Trend	0.0088	0.0060	1.4587
ecm (-1)	-0.3950	0.1088	-3.6299***
R-Squared: 0.77730 R-Bar Squared: 0.71657		Akaike Info. Criterion: 91.9537	
SER: 0.0091591		Schwarz Bayesian Criterion: 87.1682	
DW-Statistic: 2.0285		F-Statistic: 15.3577 (0.000)	
Sweden			
ARDL (2,4,3) selected based on Akaike Information Criterion			
Δ GDP	-0.2518	0.1518	-1.6586
Δ Cap. Stock	5.4099	0.5040	10.7349***
Δ Cap. Stock1	-2.9660	1.1375	-2.6074**
Δ Cap. Stock2	0.0703	1.2894	0.0545
Δ Cap. Stock3	2.9246	0.7745	0.0545***
Δ Sci. Papers	0.0704	0.0671	3.1940***
Δ Sci. Papers1	0.0704	0.0647	1.0877
Δ Sci. Papers2	0.0972	0.0595	1.6334
Δ Crisis	-0.0254	0.0082	-3.0984***
Δ Trend	0.0247	0.0035	7.0507***
ecm(-1)	-0.4025	0.0916	-4.3955***
R-Squared: 0.96989 R-Bar Squared: 0.93978		Akaike Info. Criterion: 94.5187	
SER: 0.0062654		Schwarz Bayesian Criterion: 85.4478	
DW-Statistic: 2.7216		F-Statistic: 38.0677 (0.000)	

Note: ***, **, * denotes the significance at 1%, 5% and 10% levels respectively.

Source: Author's estimations.

In summary, the results of this empirical analysis demonstrate that there is a long-run relationship between scientific publications and the level of GDP per

capita only in France, the Netherlands and Sweden out of 15 European Union countries over the period 1981-2011. While scientific publications do not have a statistically significant effect on GDP per capita in the Netherlands both in the short-run and the long-run this variable influences the level of GDP per capita in France and Sweden. Although the impact of scientific publications on the GDP per capita is negative in France in the short-run this impact becomes positive in the long-run. In Sweden scientific publications have a strong positive effect on GDP per capita both in the short-run and the long-run. These results indicate that whilst a long time is necessary to transform the scientific products into commercial products in France this transformation can be achieved even in the short-run in Sweden.

5 Conclusion

Since the beginning of 1990s research and development activities and technology have been accepted as one of the main drivers of productivity and economic growth both by economists and policy makers. Hence, the effect of these variables on the level of productivity and GDP has been investigated intensively in recent years. Although there are many studies which use input indicators such as research and development expenditures while investigating the relationship between technology and the level of productivity or GDP the number of studies which employ output indicators such as scientific publications is very low. In this empirical analysis, it is tried to fill in this gap in the existing literature.

In order to elucidate the effect of scientific publications on the level of GDP per capita in the European Union countries ARDL methodology to cointegration is applied for the 15 European Union countries over the period 1981-2011. The results of the empirical analysis indicate that there is a long-run relationship between scientific publications and the level of GDP per capita only in France, the Netherlands and Sweden out of 15 European Union countries. While in France and Sweden scientific publications have a statistically significant effect on the level of GDP per capita both in the short-run and the long-run this variable does not have a statistically significant impact on the level of GDP per capita in the Netherlands. Moreover, although the effect of scientific publications on the level of GDP per capita is negative in France in the short-run this negative effect turns out to be positive in the long-run. As for Sweden scientific publications influence the level of GDP per capita positively both in the short-run and the long-run. According to these results, it is suggested that a long time is necessary in order to transform scientific products into commercial and marketable products in France. However, in Sweden scientific products can be turned into commercial products even in the short-run.

To conclude, it is argued that scientific products do not have an effect on the level of GDP per capita in most of the European Union countries except France and Sweden. This result indicates that the cooperation between laboratories, universities and the industry is still weak in the European Union. Thus, in order to improve this cooperation and in this way to increase the contribution of scientific products to the level of economic activity science policies which empower the linkage between laboratories, universities and the industry should be implemented immediately.

References

- [1] AMECO (Annual Macro-economic Database of the European Commission's Directorate General for Economic and Financial Affairs), Economic Databases and Indicators-AMECO, 2015, http://ec.europa.eu/economy_finance/ameco/user/serie/SelectSerie.cfm.
- [2] R. L. Brown, J. Durbin and J. M. Evans, "Techniques for Testing the Constancy of Regression Relationships over Time", *Journal of the Royal Statistical Society*, **37**(2), (1975), 149-192.
- [3] J. Clemente, A. Montanes and M. Reyes, "Testing for a Unit Root in Variables with a Double Change in the Mean", *Economics Letters*, **59**(2), (1998), 175-182.
- [4] G. Elliott, T. J. Rothernberg and J. H. Stock, "Efficient Tests for an Autoregressive Unit Root", *Econometrica*, **64**(4), (1996), 813-836.
- [5] R. F. Engle and C. W. J. Granger, "Co-integration and Error Correction: Representation, Estimation and Testing", *Econometrica*, **55**(2), (1987), 251-276.
- [6] S. Ghatak and J. U. Siddiki, "The Use of the ARDL Approach in Estimating Virtual Exchange Rates in India", *Journal of Applied Statistics*, **28**(5), (2001), 573-583.
- [7] S. Ghosh, "Import Demand of Crude Oil and Economic Growth: Evidence from India", *Energy Policy*, **37**(2), (2009), 699-702.
- [8] R. Inglesi-Lotz and A. Pouris, "The Influence of Scientific Research Output of Academics on Economic Growth in South Africa: An Autoregressive Distributed Lag (ARDL) Application", *Scientometrics*, **95**(1), (2013), 129-139.
- [9] R. Inglesi-Lotz, T. Chang and R. Gupta, "Causality between Research Output and Economic Growth in BRICS", *Quality & Quantity*, **49**(1), (2015), 167-176.
- [10] A. Jalil and S. F. Mahmud, "Environment Kuznets Curve for CO₂ Emissions: A Cointegration Analysis for China", *Energy Policy*, **37**(2), (2009), 5167-5172.

- [11] S. Johansen, “Estimation and Hypothesis Testing of Cointegration Vectors in Gaussian Autoregressive Models”, *Econometrica*, **59**(6), (1991), 1551-1580.
- [12] S. Johansen, *Likelihood-Based Inference in Cointegrated Vector Auto-Regressive Models*, Oxford University Press, Oxford, 1995.
- [13] S. Johansen and K. Juselius, “Maximum Likelihood Estimation and Inference on Cointegration-with Applications to the Demand for Money”, *Oxford Bulletin of Economics and Statistics*, **52**(2), (1990), 169-210.
- [14] Y. K. Kim and K. Lee, “Different Impacts of Scientific and Technological Knowledge on Economic Growth: Contrasting Science and Technology Policy in East Asia and Latin America”, *Asian Economic Policy Review*, **10**(1), (2015), 43-66.
- [15] D. A. King, “The Scientific Impacts of Nations”, *Nature*, **430**(6997), (2004), 311-316.
- [16] R. E. Lucas, “On the Mechanics of Economic Development”, *Journal of Monetary Economics*, **22**(1), (1988), 3-42.
- [17] R. R. Kumar, P. J. Stauvermann and A. Patel, “Exploring the Link between Research and Economic Growth: An Empirical Study of China and USA”, *Quality and Quantity*, coming soon.
- [18] J. Laurenceson and J. C. H. Chai, *Financial Reform and Economic Development in China*, Edward Elgar Publishing, Cheltenham, 2003.
- [19] L. C. Lee, P. H. Lin, Y. W. Chuang and Y. Y. Lee, “Research Output and Economic Productivity: A Granger Casualty Test”, *Scientometrics*, **89**(2), (2011), 465-478.
- [20] S. A. Meo, A. A. Al Masri, A. M. Usmani, A. N. Memon and S. Z. Zaidi, “Impact of GDP, Spending on R&D, Number of Universities and Scientific Journals on Research Publications among Asian Countries”, *PLOS ONE*, **8**(6), (2013),
<http://www.plosone.org/article/fetchObject.action?uri=info:doi/10.1371/journal.pone.0066449&representation=PDF>, 1-8.
- [21] D. Moya-Anegon and V. Herrero-Solana, “Science in America-Latina: A Comparison of Bibliometric and Scientific-Technical Indicators”, *Scientometrics*, **46**(2), (1999), 299-320.
- [22] P. K. Narayan, “The Saving and Investment Nexus for China: Evidence from Cointegration Tests”, *Applied Economics*, **37**(17), (2005), 1979-1990.
- [23] S. Ng and P. Perron, “Lag Length Selection and the Construction of Unit Root Tests with Good Size and Power”, *Econometrica*, **69**(6), (2001), 1519-1554.
- [24] B. Ouattara, “The Impact of Project Aid and Programme Aid Inflows on Domestic Savings: A Case Study of Cote d’Ivoire”, *Centre for the Study of African Economies Conference on Growth, Poverty Reduction and Human Development in Africa*, (21-22 March 2004), <http://www.csae.ox.ac.uk/conferences/2004-gprahdia/papers/1m-ouattar-csae2004.pdf>.
- [25] P. Perron, “The Great Crash, the Oil Price Shock, and the Unit Root Hypothesis”, *Econometrica*, **57**(6), (1989), 1361-1401.

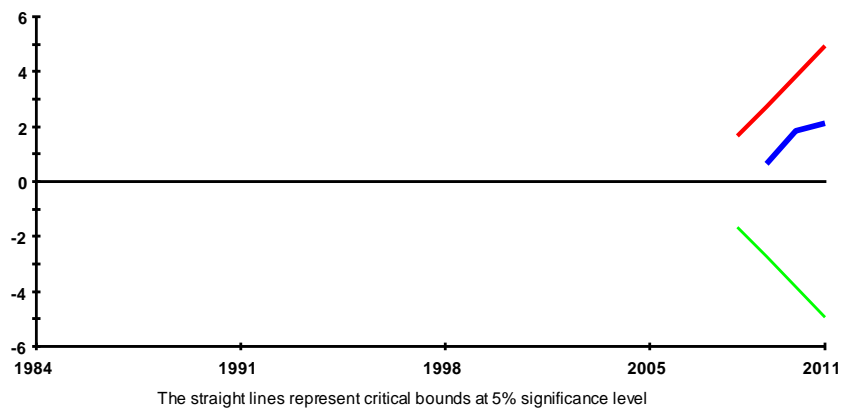
- [26] P. Perron and S. Ng, "Useful Modifications to Some Unit Root Tests with Dependent Errors and their Local Asymptotic Properties", *The Review of Economic Studies*, **63**(3), (1996), 435-463.
- [27] M. H. Pesaran, "The Role of Economic Theory in Modelling the Long Run", *The Economic Journal*, **107**(440), (1997), 178-191.
- [28] B. Pesaran and M. H. Pesaran, *Time Series Econometrics using Microfit 5.0*, Oxford University Press, Oxford, 2009.
- [29] M. H. Pesaran and R. P. Smith, "Structural Analysis of Cointegrating VARs", *Journal of Economic Surveys*, **12**(5), (1998), 471-505.
- [30] M. H. Pesaran and Y. Shin, "An Autoregressive Distributed-Lag Modelling Approach to Cointegration Analysis", *Econometrics and Economic Theory in the 20th Century: The Ragnar Frisch Centennial Symposium*, Editor: S. Strom, Cambridge University Press, Cambridge, 1999, 371-413.
- [31] M. H. Pesaran, Y. Shin and R. J. Smith, "Bounds Testing Approaches to the Analysis of Level Relationships", *Journal of Applied Econometrics*, **16**(3), (2001), 289-326.
- [32] P. C. B. Phillips and P. Perron, "Testing for a Unit Root in Time Series Regression", *Biometrika*, **75**(2), (1988), 335-346.
- [33] B. B. Rao, "Estimates of the Steady State Growth Rates for Selected Asian Countries with an Extended Solow Model", *Economic Modelling*, **27**(1), (2010), 46-53.
- [34] P. M. Romer, "Increasing Returns and Long-Run Growth", *Journal of Political Economy*, **94**(5), (1986), 1002-1037.
- [35] M. B. Shrestha and K. Chowdhury, "ARDL Modelling Approach to Testing the Financial Liberalisation Hypothesis", *University of Wollongong-Economics Working Paper Series*, **W05-15**, Department of Economics, University of Wollongong, (2005), <http://ro.uow.edu.au/cgi/viewcontent.cgi?article=1122&context=commwkpapers>.
- [36] R. M. Solow, "A Contribution to the Theory of Economic Growth", *The Quarterly Journal of Economics*, **70**(1), (1956), 65-94.
- [37] J. E. Sturm, G. H. Kuper and J. De Haan, "Modelling Government Investment and Economic Growth at the Macro Level: A Review", *Market Behaviour and Macroeconomic Modelling*, Editors: S. Kuipers, S. Brakman, H. Van Ees, Palgrave Macmillan, UK, 1998, 359-406.
- [38] H. Y. Toda and T. Yamamoto, "Statistical Inference in Vector Autoregressions with Possibly Integrated Processes", *Journal of Econometrics*, **66**(1-2), (1995), 225-250.
- [39] P. Vinkler, "Correlation between the Structure of Scientific Research, Scientometric Indicators and GDP in EU and Non-EU Countries", *Scientometrics*, **74**(2), 2008, 237-254.

- [40] World Bank, World Development Indicators Database, 2015, <http://databank.worldbank.org/data/reports.aspx?source=world-development-indicators>.

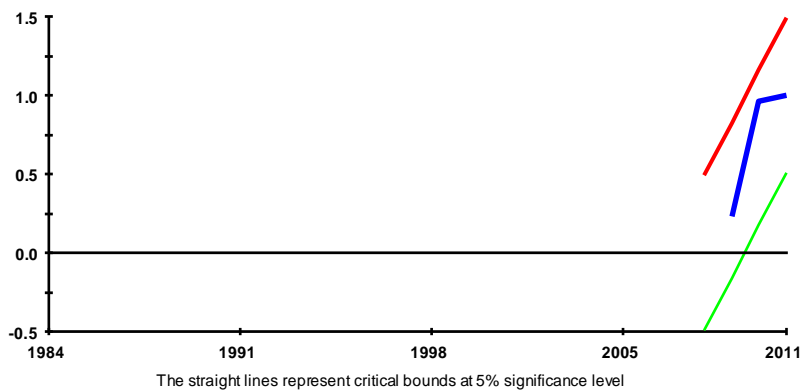
APPENDIX A

FRANCE

Plot of Cumulative Sum of Recursive Residuals



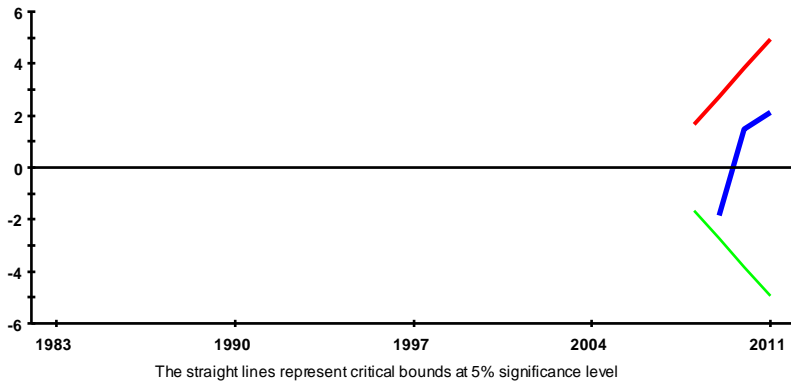
Plot of Cumulative Sum of Squares of Recursive Residuals



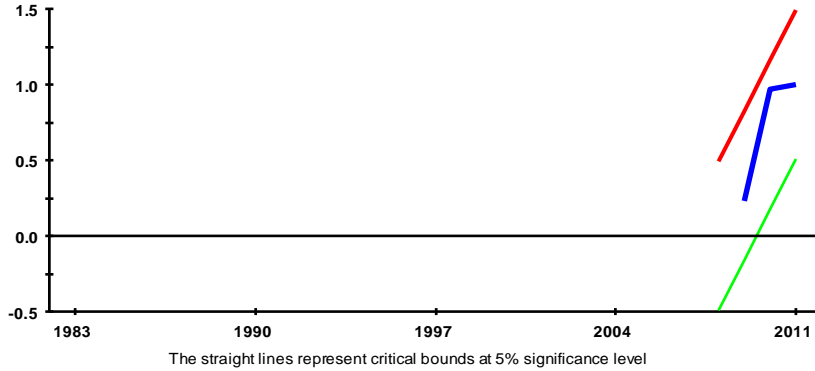
APPENDIX B

THE NETHERLANDS

Plot of Cumulative Sum of Recursive Residuals



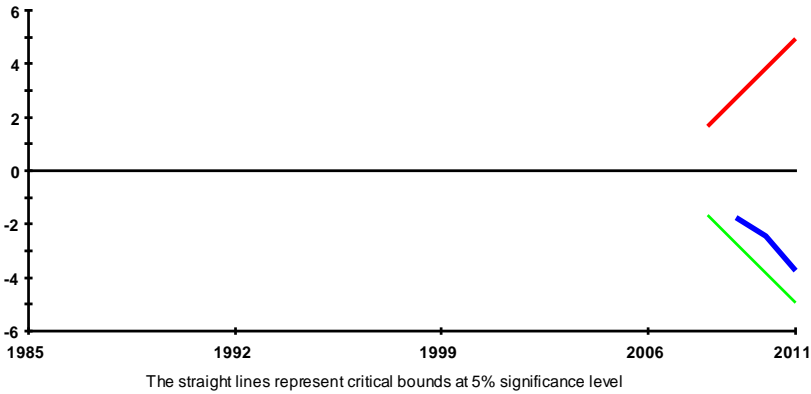
Plot of Cumulative Sum of Squares of Recursive Residuals



APPENDIX C

SWEDEN

Plot of Cumulative Sum of Recursive Residuals



Plot of Cumulative Sum of Squares of Recursive Residuals

